A chain ratio-type exponential estimator for population mean in double sampling

Nirupama Sahoo¹, Sananda Kumar Jhankar²

Abstract

In this paper we have proposed an efficient ratio-type exponential estimator for estimating the population mean of the study variable, by incorporating two auxiliary variables in twophase (double) sampling. The bias and the mean square error of the proposed estimator have been obtained up to the first order of approximation. The newly proposed estimator offers more precision in comparison to other competing estimators, theoretically as well as empirically, by considering a known value of some population parameter.

Key words: two-phase sampling, auxiliary variables, study variable, bias, mean square error, percent relative efficiency.

1. Introduction

Consider a finite population $U = (U_1, U_2, \dots, \dots, U_N)$ of N units. Let $\overline{X}, \overline{Y}$ and \overline{Z} denote the population mean, C_x , C_y and C_z denote the coefficient of variation, ρ_{yx} , ρ_{yz} and ρ_{xz} denote the correlation coefficient. Let Y be the study variable and X and Z be the auxiliary variables with corresponding value y_i, x_i, z_i ($i = 1, 2, \dots, N$). The problem is to estimate \overline{Y} in the presence of two auxiliary variable x and z.

Let $S_y^2 = \sum_{i=1}^n (y_i - \bar{Y})^2 / (N - 1)$ and $S_x^2 = \sum_{i=1}^n (x_i - \bar{X})^2 / (N - 1) S_z^2 = \sum_{i=1}^n (z_i - \bar{Z})^2 / (N - 1)$ and let $C_y = S_y / \bar{Y}$ and $C_x = S_x / \bar{X} C_z = S_z / \bar{Z}$ be the coefficients of variation of y, x and z respectively. $f_1 = \left(\frac{1}{n} - \frac{1}{N}\right) = \left(\frac{(1-f)}{n}\right)$, $f_2 = \left(\frac{1}{n'} - \frac{1}{N}\right) = \left(\frac{(1-f')}{n'}\right)$, $f_3 = f_1 - f_2 = \left(\frac{(1-f'')}{n}\right)$ where $f = \frac{n}{N}$, $f' = \frac{n'}{N}$ and $f'' = \frac{n}{n'}$ $v(\bar{y}) = f_1 \bar{Y}^2 C_y^2$

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2. Estimator with single auxiliary variable

When the population mean of the auxiliary variable *x* is not known, Sukhatme (1962) defined the two-phase sampling ratio estimator for population mean \overline{Y} as

$$t_1 = \bar{y} \left(\frac{\bar{x}'}{\bar{x}}\right) \tag{1}$$

$$MSE(t_1) = \bar{Y}^2 \Big[f_1 C_y^2 + f_3 \Big(C_x^2 - 2\rho_{yx} C_y C_x \Big) \Big]$$
(2)

The usual regression estimator in two-phase sampling is defined as

$$t_2 = \bar{y} + b_{yx}(n)(\bar{x}' - \bar{x})$$
(3)

$$MSE(t_2) = \bar{Y}^2 C_y^2 \left[f_1 \left(1 - \rho_{yx}^2 \right) + f_2 \rho_{yx}^2 \right]$$
(4)

Singh and Vishwakarma (2007) suggested exponential ratio and product type estimator for $\overline{Y}{\rm as}$

$$t_3 = \bar{y}exp\left(\frac{\bar{x}'-\bar{x}}{\bar{x}'+\bar{x}}\right) \tag{5}$$

$$t_4 = \bar{y}exp\left(\frac{x-\bar{x}'}{x+\bar{x}'}\right) \tag{6}$$

The MSEs of the estimators t_3 and t_4 respectively are

$$MSE(t_3) = \bar{Y}^2 \left[f_1 C_y^2 + \frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) \right]$$
(7)

$$MSE(t_4) = \bar{Y}^2 \left[f_1 C_y^2 + \frac{f_3}{4} \left(C_x^2 + 4\rho_{yx} C_y C_x \right) \right]$$
(8)

3. Estimator with two auxiliary variables

Chand (1975) suggested a chain ratio-type estimator for the population mean \bar{Y} defined as

$$t_5 = \bar{y}\left(\frac{\bar{x}'}{\bar{x}}\right)\left(\frac{\bar{z}}{\bar{z}'}\right); \ \bar{x} \neq 0 , \ \bar{z} \neq 0 \tag{9}$$

$$MSE(t_5) = \bar{Y}^2 \Big[f_1 C_y^2 + f_3 \Big(C_x^2 - 2\rho_{yx} C_y C_x \Big) + f_2 \Big(C_z^2 - 2\rho_{yz} C_y C_z \Big) \Big]$$
(10)

Kiregyera (1980, 1984) suggested some modification of Chand (1975) and proposed ratio to regression estimator of population mean given as

$$t_6 = \bar{y} + b_{yx}(n) \left(\frac{\bar{x}'}{\bar{z}'}\bar{Z} - \bar{x}\right) \tag{11}$$

where $b_{yx}(n)$: Sample regression coefficient of y on x based on s (Sub-sample)

 $\rho_{yx}, \rho_{yz}, \rho_{xz}$: Population correlation coefficient between the variables.

The MSE of estimator t_6 is as follows:

$$MSE(t_6) = \bar{Y}^2 C_y^2 \left[f_1 \{ 1 - \rho_{yx}^2 \} + f_2 \left\{ \rho_{yx}^2 + \rho_{yx}^2 \frac{c_z^2}{c_x^2} - 2\rho_{yx} \rho_{yz} \frac{c_z}{c_x} \right\} \right]$$
(12)

Singh and Khalid (2015) suggested the following estimator:

$$t_7 = \bar{y}exp\left(\frac{\bar{x}'(\frac{\bar{z}^*}{\bar{z}}) - \bar{x}}{\bar{x}'(\frac{\bar{z}^*}{\bar{z}}) + \bar{x}}\right)$$
(13)

where

$$\bar{z}^* = \frac{(N\bar{z} - n'\bar{z}')}{(N - n')}$$
 and $k = \frac{n'}{(N - n')}$

The required mean square error of the estimator t_7 is

$$MSE(t_7) = \bar{Y}^2 \left[f_1 C_y^2 + f_2 \left(\frac{k^2}{4} C_z^2 - k\rho_{yz} C_y C_z \right) + \frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) \right] (14)$$

Singh and Choudhury (2012) developed the following exponential chain-type ratio estimators of \overline{Y} under double sampling as

$$t_8 = \bar{y}exp\left(\frac{\left(\frac{\bar{x}'}{\bar{z}'}\right)\bar{z}-\bar{x}}{\left(\frac{\bar{x}'}{\bar{z}'}\right)\bar{z}+\bar{x}}\right)$$
(15)

The MSE of the estimator t_8 is

$$MSE(t_8) = \bar{Y}^2 \left[f_1 C_y^2 + \frac{1}{4} (f_3 C_x^2 + f_2 C_z^2) - (f_3 \rho_{yx} C_y C_x + f_2 \rho_{yz} C_y C_z) \right]$$
(16)

Motivated by Singh and Vishwakarma (2007), Yadav, Singh and Chatterjee (2013) suggested a class of chain ratio exponential type estimator for population mean \overline{Y} using information on two auxiliary variables x and z in two phase or double sampling as

$$t_9 = \bar{y} exp\left[\frac{\hat{\bar{X}}_{rd} - \bar{x}}{\hat{\bar{X}}_{rd} + \bar{x}}\right]$$
(17)

where

$$\hat{\bar{X}}_{rd} = \frac{\bar{x}'}{(a\bar{z}'+b)}(a\bar{Z}+b)$$

Thus, the above estimator can be expressed as

$$t_9 = \bar{y}exp\left[\frac{\frac{\bar{x}'}{(a\bar{z}'+b)}(a\bar{z}+b)-\bar{x}}{\frac{\bar{x}'}{(a\bar{z}'+b)}(a\bar{z}+b)+\bar{x}}\right]$$
(18)

The mean square error of the estimator t_9 is

$$MSE(t_9) = \bar{Y}^2 \left[f_1 C_y^2 + \frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) - f_2 \rho_{yz}^2 C_y^2 \right]$$
(19)

4. The suggested class of estimator

Following the previously discussed estimation procedures for two-phase sampling, we have proposed an efficient ratio-type exponential estimator

$$t_{10} = \bar{y}exp\left[k_1\left\{\frac{\bar{x}'-\bar{x}}{\bar{x}'+\bar{x}}\right\} + k_2\left\{\frac{\bar{z}'-\bar{z}}{\bar{z}'+\bar{z}}\right\}\right]$$
(20)

Where k_1 and k_2 are unknown constants. The values of k_1 and k_2 can be determined by the principle of optimality conditions.

To obtain the mean square error of the proposed estimator t_{10} , we consider

$$\bar{y} = \bar{Y}(1+e_0), \bar{x} = \bar{X}(1+e_1), \quad \bar{x}' = \bar{X}(1+e_2), \quad \bar{z}' = \bar{Z}(1+e_3)$$

Such that

$$E(e_0) = E(e_1) = E(e_2) = E(e_3) = 0$$

$$E(e_0^2) = \left(\frac{1-f}{n}\right)C_y^2, \quad E(e_1^2) = \left(\frac{1-f}{n}\right)C_x^2$$

$$E(e_2^2) = \left(\frac{1-f'}{n'}\right)C_x^2, \quad E(e_3^2) = \left(\frac{1-f'}{n'}\right)C_z^2$$

$$E(e_0e_1) = \left(\frac{1-f}{n}\right)\rho_{yx}C_yC_x, \quad E(e_0e_2) = \left(\frac{1-f'}{n'}\right)\rho_{yx}C_yC_x$$

$$E(e_0e_3) = \left(\frac{1-f'}{n'}\right)\rho_{yz}C_yC_z, \quad E(e_1e_2) = \left(\frac{1-f'}{n'}\right)C_x^2$$

$$E(e_1e_3) = \left(\frac{1-f'}{n'}\right)\rho_{xz}C_xC_z, \quad E(e_2e_3) = \left(\frac{1-f'}{n'}\right)\rho_{xz}C_xC_z$$

The mean square error of the proposed estimator t_{10} is as follows:

$$\begin{split} t_{10} &= \bar{y}exp\left[k_1\left\{\frac{e_2-e_1}{2+e_1+e_2}\right\} + k_2\left\{\frac{e_3}{2+e_3}\right\}\right] \\ &= \bar{y}exp\left[k_1\{e_2-e_1\}\frac{1}{2}\left(1-\frac{e_1}{2}-\frac{e_2}{2}+\frac{e_1^2}{4}+\frac{e_2^2}{4}\right) + k_2\left\{e_3\frac{1}{2}\left(1-\frac{e_3}{2}+\frac{e_3^2}{4}\right)\right\}\right] \\ &= \bar{y}exp\left[k_1\left\{\frac{e_2}{2}-\frac{e_1e_2}{4}-\frac{e_2^2}{4}-\frac{e_1}{2}+\frac{e_1^2}{4}+\frac{e_1e_2}{4}\right\} + k_2\left\{\frac{e_3}{2}-\frac{e_3^2}{4}\right\}\right] \\ &= \bar{Y}(1+e_0)\left[1+k_1\left\{\frac{e_2}{2}-\frac{e_1e_2}{4}-\frac{e_2^2}{4}-\frac{e_1}{2}+\frac{e_1^2}{4}+\frac{e_1e_2}{4}\right\} + k_2\left\{\frac{e_3}{2}-\frac{e_3^2}{4}\right\} + \left\{k_1\left(\frac{e_2}{2}-\frac{e_1e_2}{4}-\frac{e_2^2}{4}-\frac{e_1e_2}{4}+\frac{e_1e_2}{4}\right)\right\}^2 + \left\{k_2\left(\frac{e_3}{2}-\frac{e_3^2}{4}\right)\right\}^2\right] \end{split}$$

Neglecting the term having power greater than two, we have

$$t_{10} = \bar{Y} \left[1 + k_1 \frac{e_2}{2} - k_1 \frac{e_2^2}{4} - k_1 \frac{e_1}{2} + k_1 \frac{e_1^2}{4} + k_2 \frac{e_3}{2} - k_2 \frac{e_3^2}{4} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_1^2}{8} + k_2^2 \frac{e_3^2}{8} + e_0 + k_1 \frac{e_0 e_2}{2} - k_1 \frac{e_0 e_1}{2} + k_2 \frac{e_0 e_3}{2} \right]$$

or

$$t_{10} - \bar{Y} = \bar{Y} \left[k_1 \frac{e_2}{2} - k_1 \frac{e_2^2}{4} - k_1 \frac{e_1}{2} + k_1 \frac{e_1^2}{4} + k_2 \frac{e_3}{2} - k_2 \frac{e_3^2}{4} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_3^2}{8} + e_0 + k_1 \frac{e_0 e_2}{2} - k_1 \frac{e_0 e_1}{2} + k_2 \frac{e_0 e_3}{2} \right]$$
(21)

The bias of the estimator t_{10} can be obtained by taking expectation on both sides of equation (21) and is given by

$$B(t_{10}) = \overline{Y}^2 \left[k_1 \frac{e_1^2}{4} - k_1 \frac{e_2^2}{4} - k_2 \frac{e_3^2}{4} + k_1^2 \frac{e_2^2}{8} + k_1^2 \frac{e_1^2}{8} + k_2^2 \frac{e_3^2}{8} + k_1 \frac{e_0 e_2}{2} - k_1 \frac{e_0 e_1}{2} + k_2 \frac{e_0 e_3}{2} \right]$$
(22)

By squaring and taking expectation on both sides of equation (21), we get the mean square error of t_{10} to the first degree of approximation.

$$\begin{split} E(t_{10} - \bar{Y})^2 &= \bar{Y}^2 E\left[k_1^2 \frac{e_2^2}{4} + k_1^2 \frac{e_1^2}{4} + k_2^2 \frac{e_3^2}{4} + e_0^2 - k_1^2 \frac{e_1 e_2}{2} + k_1 k_2 \frac{e_2 e_3}{2} + k_1 e_0 e_2 - k_1 k_2 \frac{e_1 e_3}{2} - k_1 e_0 e_1 + k_2 e_0 e_3\right] \\ &= \bar{Y}^2 \left[\left(\frac{1-f}{n}\right) C_y^2 + \frac{k_1^2}{4} \left(\frac{1-f''}{n}\right) C_x^2 + \frac{k_2^2}{4} \left(\frac{1-f'}{n'}\right) C_z^2 - k_1 \left(\frac{1-f''}{n}\right) \rho_{yx} C_y C_x + k_2 \left(\frac{1-f'}{n'}\right) \rho_{yz} C_y C_z \right] \\ MSE(t_{10}) &= \bar{Y}^2 \left[f_1 C_y^2 + \frac{k_1^2}{4} f_3 C_x^2 + \frac{k_2^2}{4} f_2 C_z^2 - k_1 f_3 \rho_{yx} C_y C_x + k_2 f_2 \rho_{yz} C_y C_z \right] \end{split}$$
(23)

Now, we have to find out the optimum values of k_1 and k_2

$$\frac{\partial MSE(t_{10})}{\partial k_1} = 0$$

$$\Rightarrow k_{1_{opt}} = \frac{2\rho_{yx}C_yC_x}{C_x^2}$$

$$\Rightarrow k_{1_{opt}} = 2\rho_{yx}\left(\frac{C_y}{C_x}\right)$$

Now, for k_2

$$\frac{\partial MSE(t_{10})}{\partial k_2} = 0$$

$$\Rightarrow k_{2opt} = \frac{2\rho_{yz}C_yC_z}{C_z^2}$$

$$\Rightarrow k_{2opt} = -2\rho_{yz}\left(\frac{C_y}{C_z}\right)$$

By substituting above k_{1opt} and k_{2opt} in equation (23) we obtained the minimum mean square error of the estimator t_{10} as

$$MSE(t_{10})_{opt} = \bar{Y}^2 \left[f_1 C_y^2 + \left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right]$$

$$(24)$$

5. Theoretical Comparison

The proposed estimator t_{10} under its optimality condition is more efficient than the existing estimators t_i (i = 1, 2, ..., 9) if and only if the following conditions hold.

(i)
$$MSE(t_1) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[f_3 \left(C_x^2 - 2\rho_{yx} C_y C_x \right) \right]$$

$$(25)$$

(ii)
$$MSE(t_2) - MSE(t_{10})_{opt} > 0$$

$$\begin{bmatrix} f_1 C_y^2 + \left(\rho_{yx} \left(\frac{C_y}{C_x}\right)\right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z}\right)\right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x}\right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z}\right) f_2 \rho_{yz} C_y C_z \end{bmatrix} < C_y^2 \begin{bmatrix} f_1 \left(1 - \rho_{yx}^2\right) + f_2 \rho_{yx}^2 \end{bmatrix}$$
(26)

(iii)
$$MSE(t_3) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[\frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) \right]$$

$$(27)$$

(iv) $MSE(t_4) - MSE(t_{10})_{opt} > 0$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[\frac{f_3}{4} \left(C_x^2 + 4\rho_{yx} C_y C_x \right) \right]$$

$$(28)$$

(v)
$$MSE(t_{5}) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_{y}}{C_{x}} \right) \right)^{2} f_{3}C_{x}^{2} + \left(\rho_{yz} \left(\frac{C_{y}}{C_{z}} \right) \right)^{2} f_{2}C_{z}^{2} - 2\rho_{yx} \left(\frac{C_{y}}{C_{x}} \right) f_{3}\rho_{yx}C_{y}C_{x} - 2\rho_{yz} \left(\frac{C_{y}}{C_{z}} \right) f_{2}\rho_{yz}C_{y}C_{z} \right] < \left[f_{3} \left(C_{x}^{2} - 2\rho_{yx}C_{y}C_{x} \right) + f_{2} \left(C_{z}^{2} - 2\rho_{yz}C_{y}C_{z} \right) \right]$$
(29)

(vi)
$$MSE(t_{6}) - MSE(t_{10})_{opt} > 0$$

$$\left[f_{1}C_{y}^{2} + \left(\rho_{yx} \left(\frac{C_{y}}{C_{x}} \right) \right)^{2} f_{3}C_{x}^{2} + \left(\rho_{yz} \left(\frac{C_{y}}{C_{z}} \right) \right)^{2} f_{2}C_{z}^{2} - 2\rho_{yx} \left(\frac{C_{y}}{C_{x}} \right) f_{3}\rho_{yx}C_{y}C_{x} - 2\rho_{yz} \left(\frac{C_{y}}{C_{z}} \right) f_{2}\rho_{yz}C_{y}C_{z} \right] < C_{y}^{2} \left[f_{1}\{1 - \rho_{yx}^{2}\} + f_{2} \left\{ \rho_{yx}^{2} + \rho_{yx}^{2} \frac{C_{z}^{2}}{C_{x}^{2}} - 2\rho_{yx}\rho_{yz} \frac{C_{z}}{C_{x}} \right\} \right]$$
(30)

(vii)
$$MSE(t_7) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[f_2 \left(\frac{k^2}{4} C_z^2 - k\rho_{yz} C_y C_z \right) + \frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) \right]$$
(31)

(viii)
$$MSE(t_8) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[\frac{1}{4} (f_3 C_x^2 + f_2 C_z^2) - (f_3 \rho_{yx} C_y C_x + f_2 \rho_{yz} C_y C_z) \right]$$
(32)

(ix)
$$MSE(t_{9}) - MSE(t_{10})_{opt} > 0$$

$$\left[\left(\rho_{yx} \left(\frac{C_y}{C_x} \right) \right)^2 f_3 C_x^2 + \left(\rho_{yz} \left(\frac{C_y}{C_z} \right) \right)^2 f_2 C_z^2 - 2\rho_{yx} \left(\frac{C_y}{C_x} \right) f_3 \rho_{yx} C_y C_x - 2\rho_{yz} \left(\frac{C_y}{C_z} \right) f_2 \rho_{yz} C_y C_z \right] < \left[\frac{f_3}{4} \left(C_x^2 - 4\rho_{yx} C_y C_x \right) - f_2 \rho_{yz}^2 C_y^2 \right]$$
(33)

6. Empirical Study

To judge the superiority of our newly proposed estimator over the competing estimators we have taken the following numerical values of different population parameters from two different population data sets.

Population-1 [Source: Singh (1967)]

The variables are

- y: Number of females employed
- *x*: Number of females in services
- z: Number of educated females

$$\begin{split} N &= 61n' = 20n = 10\bar{Y} = 7.46\\ \bar{X} &= 5.31\bar{Z} = 179.00C_y = 0.7103C_x = 0.7587\\ C_z &= 0.2515\rho_{yx} = 0.7737\rho_{yz} = -0.2070\rho_{xz} = -0.0033 \end{split}$$

Population-2 [Source: Murthy (1967) pp. 226]

The variables are

- y: Output
- x: Number of workers
- z: Fixed capital

$$N = 61n' = 20n = 10\overline{Y} = 7.46$$

$$\overline{X} = 5.31\overline{Z} = 179.00C_y = 0.7103C_x = 0.7587$$

$$C_z = 0.2515\rho_{yx} = 0.7737\rho_{yz} = -0.2070\rho_{xz} = -0.0033$$

To compare the efficiency of the proposed estimator and the considered existing estimators, we have computed the percent relative efficiencies (PREs).

The formula for calculating the percent relative efficiency is given by

$$PRE(t_i) = \left[\frac{var(\overline{y})}{MSE(t_i)}\right] \times 100; \qquad (34)$$

where *i* = 1,2,3,4,5,6,7,8,9,10

Findings are given in Table 1.

Estimators	Population-1	Population-2
t_1	144.1214	39.1380
t_2	155.8702	233.9572
t_3	147.6438	180.6875
t_4	60.1505	25.6294
t_5	124.5640	36.8749
t_6	140.9066	607.6388
t ₇	144.1214	285.2960
t_8	139.3606	361.375
t9	151.3525	373.8362
t ₁₀	159.9441	677.5781

Table 1:

7. Conclusion

On account of the percent relative efficiencies of the estimators as shown in Table 1, it has been observed that the performance of the proposed estimator is better than the other existing estimators. Hence, it is recommended for use in practice.

Acknowledgements

We are very grateful to the referees and members of the board of editors for their useful suggestions.

References

- Anderson, T. W., (1958). An Introduction to Multivariate Statistical Analysis. John Wiley& Sons, Inc., New York.
- Bahl, S., Tuteja, R. K., (1991). Ratio and product type exponential estimators, *Journal* of information and optimization sciences, 12(1), pp. 159–164.
- Bandyopadhyay, A., Singh, G. N., (2014). Some modified classes of chain type estimators of population mean in two phase sampling, *International Journal of Statistics and Economics*, 14(2), pp. 92–103.

- Chakraborty, R. P., (1968). Contribution to the theory of ratio-type Estimators, Ph.D. Thesis, Texas A and M University, U.S.A.
- Chand, L., (1975). Some ratio-type estimators based on two or more Auxiliary variables, Ph.D. diss., Iowa State University, Ames, Iowa.
- Cochran, W. G., (1977). Sampling Techniques, 3rd edition, John Wiley, NewYork.
- Dash, P. R., Mishra, G., (2011). An improved class of estimators in two-phase sampling using two auxiliary variables, *Communications in Statistics-Theory and Methods*, 40(24), pp. 4347–4352.
- Fisher, R. A., (1936). The use of multiple measurements in taxonomic problems, *Annals of eugenics*, 7(2), pp. 179–188.
- Gupta, P. C., (1978). On some quadratic and higher degree ratio and product estimators, *Journal of Indian Society of Agricultural Statistics*, 30, No. 2, pp. 71–80.
- Kadilar, C., Cingi, H., (2006). Ratio estimators for the population variance in simple and stratified random sampling, *Appl. Math Computation* 173(2), pp.1047–1059.
- Kiregyera, B., (1980). A chain ratio-type estimator in finite population double samplingusing two auxiliary variables, *Metrika*, 27, pp.217–223.
- Kiregyera, B., (1984). Regression-type estimators using two auxiliary variables and the model of double sampling from finite populations, *Metrika*, 31, pp. 215–226.
- Kumar, K., Kumar, M., (2017). Two Phase Sampling Exponential Type Estimators for Ratio and Product of two Population Means in the Presence of Non-response, *International Journal of Scientific Research in Mathematical and Statistical Sciences*, Vol. 4, Issue.6, pp. 26–34.
- Mukherjee, R., Rao, T. J., and Vijayan, K., (1987). Regression type estimators using multiple auxiliary information, *Australian Journal of Statistics*, 29(3), pp. 244–254
- Murthy, M. N., (1967). Sampling Theory and Methods, Statistical Publishing Society, Calcutta, India.
- Panda, K. B., Sahoo, N., (2017). Estimation of Finite Population Mean Using Variable Transformation, *Int. J. Agricult. Stat. Sci.*, Vol. 13, No. 1, pp. 141–144.
- Prasad, B., Singh, R. S., and Singh, H. P., (1996). Some chain ratio type estimators for ratio of two population means using two auxiliary characters in two phase sampling, *Metron*, *LIV*, No. 1-2, pp. 95–113.
- Reddy, V. N., (1973). On ratio and product methods of estimation, *Sankhyā: The Indian Journal of Statistics*, Series B, pp. 307–316.

- Reddy, V. N., (1978). A study on the use of prior knowledge on certain population parameters in estimation, *Sankhya C*, 40, pp. 29–37.
- Singh, B. K., Choudhury, S., (2012). Exponential chain ratio and product type estimators for finite population mean under double sampling scheme. *Journal of Science Frontier Research in Mathematics and Design Sciences*, 12(6), pp. 0975– 5896.
- Singh, G. N., (2001). On the use of transformed auxiliary variable in the estimation of population mean in two-phase sampling, *Statistics in Transition*, 5(3), pp. 405–416.
- Singh, G. N., Khalid, M., (2015). Exponential chain dual to ratio and regression type estimators of population mean in two-phase sampling, *Statistica*, 75(4), pp. 379– 389.
- Singh, M.P., (1967). Ratio-cum-product method of estimation, Metrika, 12, pp. 34-42.
- Sing, G. N., Sharma, A. K., (2014). An improved estimation procedure of population mean in two phase sampling, *Journal of International Academy of Physical Sciences*, 18(1), pp. 1–10.
- Singh, G. N., Sharma, A. K., (2015). Some improved estimators for population mean in double sampling, *International Journal of Mathematics and Computation*, 26(3), pp. 36–45.
- Singh, H. P., Gangele, R. K., (1999). Classes of almost unbiased ratio and product-type estimators in two phase sampling, *Statistica*, LIX, No. 1, pp. 109–124.
- Singh, H. P., Tailor, R., (2005a). Estimation of finite population mean with known coefficient of variation of an auxiliary character, *Statistica*, LXV, No. 3, pp. 301– 313.
- Singh, H. P., Tailor, R., (2005b). Estimation of finite population mean using known correlation coefficient between auxiliary characters, *Statistica*, LXV, No. 4, pp. 407– 418.
- Singh, H. P., Singh, S., and Kim, J. M. (2006). General families of chain ratio type estimators of the population mean with known coefficient of variation of the second auxiliary variable in two phase sampling, *Journal of the Korean Statistical Society*, 35, No. 4, pp. 377–395.
- Singh, H. P., Vishwakarma, G. K., (2007). Modified exponential ratio and product estimators for finite population mean in double sampling, *Austrian Journal of Statistics*, 36, No. 3, pp. 217–225.

- Singh, H. P., Mathur, N., and Chandra, P., (2009). A chain-type estimator for population variance using two auxiliary variables in two-phase sampling. *Statistics in Transition-new series*, 10, No. 1, pp. 75–84.
- Singh, H. P., Vishwakarma, G. K., (2007). Modified exponential ratio and product estimators for finite population mean in double sampling, *Austrian journal of statistics*, 36(3), pp. 217–225.
- Singh, G. N., Pandey, A. K., and Jaiswal, A. K., (2019). An efficient class of chain-type exponential estimators for population mean under two phase sampling scheme, *Journal of reliability and statistical studies*, Vol. 12, Issue 1(2019), pp. 23–40.
- Singh, V. K., Shukla, D., (1987). One parameter family of factor-type ratio estimators, *Metron*, 45(1-2), pp. 273–283.
- Singh, V. K., Singh, G. N., (1991). Chain type regression estimators with two auxiliary variables under double sampling scheme, *Metron*, 49, pp. 279–289.
- Srivenkataramana, T., Tracy, D. S., (1980). An alternative to ratio method in sample surveys, *Annals of the Institute of Statistical Mathematics*, 32(1), pp. 111–120.
- Srivenkataramana, T., Tracy, D. S., (1989). Two phase sampling for selection with probability proportional to size in sample surveys, *Biometrika*, 76, pp. 818–821.
- Shukla, G. K., (1966). An alternative multivariate ratio estimate for finite population, *Calcutta Statistical Association Bulletin*, 15(2–3), pp. 127–134.
- Upadhyaya, L. N., Kushwaha, K. S., and Singh, H. P., (1990). A modified chain ratio type estimator in two-phase sampling using multi-auxiliary information, *Metron*, 48, pp. 381–393.
- Upadhyaya, L. N., Singh, G. N., (2001). Chain type estimators using transformed auxiliary variable in two-phase sampling, *Advances in Modeling and Analysis*, 38, No. 1-2,pp. 1–10.
- Upadhyaya, L. N., Singh, H. P., and Singh, S., (2004). A family of almost unbiased estimators for negatively correlated variables using Jackknife Technique, *Statistica*, LXIV, 4, pp.767–778.
- Upadhyaya, L. N., Singh, H. P., and Tailor, R., (2006). Estimation of mean with known coefficient of variation of an auxiliary variable in two phase sample, *Statistics in Transition*, 7, No. 6, pp. 1327–1344.
- Vos, J. W. E., (1980). Mixing of direct, ratio and product method estimators, *Statist. Neerl.*, 34, pp. 209–218.

- Walsh, J. E., (1970). Generalization of ratio estimate for population total, *Sankhya*, *A*, 32, pp. 99–106.
- Yadav, R., Upadhyaya L. N., Singh, H.P., and Chatterjee, S., (2013). A chain ratio exponential type estimator in two phase sampling using auxiliary information, *Statistica*, anno LXXIII, No. 2, pp. 221–234.